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## Statistics Problems for the NIKHEF Onderzoekschool Subatomaire Fysica (part 2)

**Exercise 4:** Consider a likelihood  $L(x|\theta)$ , which gives the probability for the data x given a parameter  $\theta$ . The Jeffreys prior for  $\theta$  is given by

$$\pi(\theta) \propto \sqrt{I(\theta)}$$
, (1)

where

$$I(\theta) = -E\left[\frac{\partial^2 \ln L}{\partial \theta^2}\right] \tag{2}$$

is the expected Fisher Information.

(a) Show that

$$-E\left[\frac{\partial^2 \ln L}{\partial \theta^2}\right] = E\left[\left(\frac{\partial \ln L}{\partial \theta}\right)^2\right],\tag{3}$$

providing that the set of the allowed values of x does not depend on  $\theta$ .

Hint: Write the right-hand side of (3) as

$$E\left[\left(\frac{\partial\ln L}{\partial\theta}\right)^2\right] = \int \left[\frac{\partial}{\partial\theta}\left(L\frac{\partial\ln L}{\partial\theta}\right) - L\frac{\partial^2\ln L}{\partial\theta^2}\right]dx \tag{4}$$

and use the fact that one can bring the derivative  $\partial/\partial\theta$  outside of the integral as long as the region of integration does not depend on  $\theta$ . Also use the fact the integral of  $L(x|\theta)$  over all x is equal to unity for any  $\theta$ .

(b) Suppose one uses the Jeffreys prior for  $\theta$  to obtain the posterior pdf

$$p(\theta|x) \propto L(x|\theta)\pi(\theta)$$
, (5)

Suppose now that one transforms to a new parameter  $\eta(\theta)$ , such that the posterior pdf for  $\eta$  is

$$p(\eta|x) = p(\theta|x) \left| \frac{\partial \theta}{\partial \eta} \right| .$$
(6)

Show that one arrives at the same posterior pdf as (6) by beginning directly from the Jeffreys prior for  $\eta$ . That is, inference made using the Jeffreys prior is invariant under a parameter transformation.

**Exercise 5:** Consider an experiment where one measures a number of events n, which is modeled as following a Poisson distribution with mean s + b, where s and b are the contributions from signal and background processes, respectively. To constrain the parameter b, one carries out a control measurement that counts a number of events m, which follows a Poisson distribution with mean  $\tau b$ , where  $\tau$  is a known scale factor.

The problem thus contains a single parameter of interest, s, and a nuisance parameter b. The likelihood can be written

$$L(n,m|s,b) = \frac{(s+b)^n}{n!} e^{-(s+b)} \frac{(\tau b)^m}{m!} e^{-\tau b} , \qquad (7)$$

Show that the Maximum Likelihood (ML) estimators for s and b are

$$\hat{s} = n - m/\tau , \qquad (8)$$

$$\hat{b} = m/\tau , \qquad (9)$$

and that the conditional ML estimator for b given s is

$$\hat{\hat{b}}(s) = \frac{m+n-(1+\tau)s + \sqrt{(m+n-(1+\tau)s)^2 + 4(1+\tau)ms}}{2(1+\tau)} \,. \tag{10}$$

The quantities  $\hat{s}$ ,  $\hat{b}$  and  $\hat{b}$  are what we require to compute the profile likelihood ratio

$$\lambda(s) = \frac{L(s,\hat{b})}{L(\hat{s},\hat{b})} \,. \tag{11}$$

(b) Write a Monte Carlo program that generates values of n and m according to hypothesized values of s and b (e.g., use b = 20,  $\tau = 1$  and s = 0) and use these to evaluate the profile likelihood ratio. (See the program runSigCalc\_MC.cc on the course webpage for examples of how to generate Poisson distributed values by using the ROOT class TRandom3.)

To carry out a test of s = 0, we can use the statistic

$$q_0 = \begin{cases} -2\ln\lambda(0) & \hat{s} \ge 0 ,\\ 0 & \hat{s} < 0 . \end{cases}$$
(12)

Generate the distribution of  $q_0$  assuming s = 0, b = 20 and  $\tau = 1$ . In the large sample limit, this should approach a "half-chi-square" distribution for one degree of freedom (a delta function at zero plus a chi-square distribution, each with a weight of one half). Check to what extent this holds for different values of b (e.g., b = 2, 20, 200).

(c) Extend the Monte Carlo program from (b) to compute the distribution of  $q_0$  by generating data with a nonzero value of s, e.g., take b = 20,  $\tau = 1$ , s = 10. Find the median value of  $q_0$  under assumption of this value of s, and thus find the median discovery significance for this s (i.e., the discovery sensitivity). Compare this to the value based on the Asimov data set,

$$\operatorname{med}[Z_0|s,b] \approx \sqrt{2\left((s+b)\ln(1+s/b)-s\right)}$$
 (13)