Expected discovery significance for counting experiment with background uncertainty



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Outline / executive summary

I. Discovery sensitivity for counting experiment with *b* known:

(a)
$$\frac{s}{\sqrt{b}}$$

(b) Profile likelihood ratio test & Asimov:

$$\sqrt{2\left((s+b)\ln\left(1+\frac{s}{b}\right)-s\right)}$$

II. Discovery sensitivity with uncertainty in b, σ_b :

(a)
$$\frac{s}{\sqrt{b+\sigma_b^2}}$$

(b) Profile likelihood ratio test & Asimov:

$$\left[2\left((s+b)\ln\left[\frac{(s+b)(b+\sigma_b^2)}{b^2+(s+b)\sigma_b^2}\right] - \frac{b^2}{\sigma_b^2}\ln\left[1 + \frac{\sigma_b^2s}{b(b+\sigma_b^2)}\right]\right)\right]^{1/2}$$

Counting experiment with known background Count a number of events $n \sim Poisson(s+b)$, where s = expected number of events from signal, b = expected number of background events.

To test for discovery of signal compute *p*-value of *s*=0 hypothesis,

$$p = P(n \ge n_{\text{obs}}|b) = \sum_{n=n_{\text{obs}}}^{\infty} \frac{b^n}{n!} e^{-b} = 1 - F_{\chi^2}(2b; 2n_{\text{obs}})$$

Usually convert to equivalent significance: $Z = \Phi^{-1}(1-p)$ where Φ is the standard Gaussian cumulative distribution, e.g., Z > 5 (a 5 sigma effect) means $p < 2.9 \times 10^{-7}$.

To characterize sensitivity to discovery, give expected (mean or median) Z under assumption of a given s.

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 s/\sqrt{b} for expected discovery significance For large s + b, $n \to x \sim \text{Gaussian}(\mu, \sigma)$, $\mu = s + b$, $\sigma = \sqrt{(s + b)}$. For observed value x_{obs} , *p*-value of s = 0 is $\text{Prob}(x > x_{\text{obs}} | s = 0)$,:

$$p_0 = 1 - \Phi\left(\frac{x_{\rm obs} - b}{\sqrt{b}}\right)$$

Significance for rejecting s = 0 is therefore

$$Z_0 = \Phi^{-1}(1 - p_0) = \frac{x_{\text{obs}} - b}{\sqrt{b}}$$

Expected (median) significance assuming signal rate s is

$$\mathrm{median}[Z_0|s+b] = \frac{s}{\sqrt{b}}$$

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Better approximation for significance Poisson likelihood for parameter *s* is

> $L(s) = \frac{(s+b)^n}{n!} e^{-(s+b)}$ For now no nuisance

To test for discovery use profile likelihood ratio:

$$q_0 = \begin{cases} -2\ln\lambda(0) & \hat{s} \ge 0 \ , \\ 0 & \hat{s} < 0 \ . \end{cases} \qquad \lambda(s) = \frac{L(s, \hat{\hat{\theta}}(s))}{L(\hat{s}, \hat{\theta})}$$

So the likelihood ratio statistic for testing s = 0 is

$$q_0 = -2\ln\frac{L(0)}{L(\hat{s})} = 2\left(n\ln\frac{n}{b} + b - n\right) \quad \text{for } n > b, \ 0 \text{ otherwise}$$

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params.

Approximate Poisson significance (continued)

For sufficiently large s + b, (use Wilks' theorem),

$$Z = \sqrt{2\left(n\ln\frac{n}{b} + b - n\right)} \quad \text{for } n > b \text{ and } Z = 0 \text{ otherwise.}$$

To find median[*Z*|*s*], let $n \rightarrow s + b$ (i.e., the Asimov data set):

$$Z_{\rm A} = \sqrt{2\left(\left(s+b\right)\ln\left(1+\frac{s}{b}\right) - s\right)}$$

This reduces to s/\sqrt{b} for s << b.

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 $n \sim \text{Poisson}(s+b)$, median significance, assuming *s*, of the hypothesis s = 0

CCGV, EPJC 71 (2011) 1554, arXiv:1007.1727



"Exact" values from MC, jumps due to discrete data.

Asimov $\sqrt{q_{0,A}}$ good approx. for broad range of *s*, *b*.

 s/\sqrt{b} only good for $s \ll b$.

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Extending s/\sqrt{b} to case where b uncertain

The intuitive explanation of s/\sqrt{b} is that it compares the signal, *s*, to the standard deviation of *n* assuming no signal, \sqrt{b} .

Now suppose the value of *b* is uncertain, characterized by a standard deviation σ_b .

A reasonable guess is to replace \sqrt{b} by the quadratic sum of \sqrt{b} and σ_b , i.e.,

$$\operatorname{med}[Z|s] = \frac{s}{\sqrt{b + \sigma_b^2}}$$

This has been used to optimize some analyses e.g. where σ_b cannot be neglected.

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Profile likelihood with b uncertain

This is the well studied "on/off" problem: Cranmer 2005; Cousins, Linnemann, and Tucker 2008; Li and Ma 1983,...

Measure two Poisson distributed values:

 $n \sim \text{Poisson}(s+b)$ (primary or "search" measurement) $m \sim \text{Poisson}(\tau b)$ (control measurement, τ known)

The likelihood function is

$$L(s,b) = \frac{(s+b)^n}{n!} e^{-(s+b)} \frac{(\tau b)^m}{m!} e^{-\tau b}$$

Use this to construct profile likelihood ratio (*b* is nuisance parmeter): $L(0, \hat{b}(0))$

$$\lambda(0) = \frac{L(0, b(0))}{L(\hat{s}, \hat{b})}$$

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Asymptotic significance

Use profile likelihood ratio for q_0 , and then from this get discovery significance using asymptotic approximation (Wilks' theorem):

$$Z = \sqrt{q_0}$$
$$= \left[-2\left(n\ln\left[\frac{n+m}{(1+\tau)n}\right] + m\ln\left[\frac{\tau(n+m)}{(1+\tau)m}\right]\right) \right]^{1/2}$$

for $n > \hat{b}$ and Z = 0 otherwise.

Essentially same as in:

Robert D. Cousins, James T. Linnemann and Jordan Tucker, NIM A 595 (2008) 480– 501; arXiv:physics/0702156.

Tipei Li and Yuqian Ma, Astrophysical Journal 272 (1983) 317–324.

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Asimov approximation for median significance

To get median discovery significance, replace *n*, *m* by their expectation values assuming background-plus-signal model:

$$n \to s + b$$

$$m \to \tau b$$

$$Z_{\rm A} = \left[-2\left((s+b) \ln\left[\frac{s+(1+\tau)b}{(1+\tau)(s+b)}\right] + \tau b \ln\left[1+\frac{s}{(1+\tau)b}\right] \right) \right]^{1/2}$$
Or use the variance of $\hat{b} = m/\tau$, $V[\hat{b}] \equiv \sigma_b^2 = \frac{b}{\tau}$, to eliminate τ :

$$Z_{\rm A} = \left[2\left((s+b) \ln\left[\frac{(s+b)(b+\sigma_b^2)}{b^2+(s+b)\sigma_b^2}\right] - \frac{b^2}{\sigma_b^2} \ln\left[1+\frac{\sigma_b^2 s}{b(b+\sigma_b^2)}\right] \right) \right]^{1/2}$$

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Limiting cases

Expanding the Asimov formula in powers of *s/b* and σ_b^2/b (= 1/ τ) gives

$$Z_{\rm A} = \frac{s}{\sqrt{b + \sigma_b^2}} \left(1 + \mathcal{O}(s/b) + \mathcal{O}(\sigma_b^2/b) \right)$$

So the "intuitive" formula can be justified as a limiting case of the significance from the profile likelihood ratio test evaluated with the Asimov data set.

Testing the formulae: s = 2



Structure in points (MC) is due to discreteness of data.

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Testing the formulae: s = 5



Testing the formulae: s = 10



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Summary

Simple formula for expected discovery significance based on profile likelihood ratio test and Asimov approximation:

$$Z_{\rm A} = \left[2 \left((s+b) \ln \left[\frac{(s+b)(b+\sigma_b^2)}{b^2 + (s+b)\sigma_b^2} \right] - \frac{b^2}{\sigma_b^2} \ln \left[1 + \frac{\sigma_b^2 s}{b(b+\sigma_b^2)} \right] \right) \right]^{1/2}$$

For details see accompanying note on agenda page or www.pp.rhul.ac.uk/~cowan/stat/medsig/medsigNote.pdf For large *b*, all formulae OK.

For small *b*, s/\sqrt{b} and $s/\sqrt{(b+\sigma_b^2)}$ overestimate the significance.

Could be important in optimization of searches with low background.

Formula maybe also OK if model is not simple on/off experiment, e.g., several background control measurements (check this).

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Ingredients for profile likelihood ratio

To construct the profile likelihood ratio we need the estimators:

$$\begin{aligned} \hat{s} &= n - m/\tau , \\ \hat{b} &= m/\tau , \\ \hat{b}(s) &= \frac{n + m - (1 + \tau)s + \sqrt{(n + m - (1 + \tau)s)^2 + 4(1 + \tau)sm}}{2(1 + \tau)} \end{aligned}$$

and in particular to test for discovery (s = 0),

$$\hat{\hat{b}}(0) = \frac{n+m}{1+\tau}$$

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