

## Chapter 9

# Statistical Errors, Confidence Intervals and Limits

**Exercise 9.1:** Suppose an estimator  $\hat{\theta}$  is Gaussian distributed about the parameter's true value  $\theta$  with a standard deviation  $\sigma_{\hat{\theta}}$ . Assume that  $\sigma_{\hat{\theta}}$  is known.

- (a) Sketch the functions  $u_{\alpha}(\theta)$  and  $v_{\beta}(\theta)$  defining the confidence belt (cf. SDA Section 9.2).  
(b) Show that the central confidence interval for  $\theta$  at a confidence level  $1 - \gamma$  is given by

$$[\hat{\theta} - \sigma_{\hat{\theta}}\Phi^{-1}(1 - \gamma/2), \hat{\theta} + \sigma_{\hat{\theta}}\Phi^{-1}(1 - \gamma/2)], \quad (9.1)$$

where  $\Phi^{-1}$  is the quantile of the standard Gaussian.

**Exercise 9.2:** (a) Consider  $n$  observations of an exponentially distributed variable  $x$  with mean  $\xi$ . The ML estimator for  $\xi$  (see SDA (6.6)) is given by

$$\hat{\xi} = \frac{1}{n} \sum_{i=1}^n x_i \quad (9.2)$$

and the p.d.f. for  $\hat{\xi}$  (cf. SDA equation (10.25)) is

$$g(\hat{\xi}; \xi) = \frac{n^n}{(n-1)!} \frac{\hat{\xi}^{n-1}}{\xi^n} e^{-n\hat{\xi}/\xi}. \quad (9.3)$$

- (a) Show that the curves defining the confidence belt,  $u_{\alpha}(\xi)$  and  $v_{\beta}(\xi)$ , are given by

$$\begin{aligned} u_{\alpha}(\xi) &= \frac{\xi}{2n} F_{\chi^2}^{-1}(1 - \alpha; 2n), \\ v_{\beta}(\xi) &= \frac{\xi}{2n} F_{\chi^2}^{-1}(\beta; 2n), \end{aligned} \quad (9.4)$$

where  $F_{\chi^2}^{-1}$  is the quantile of the  $\chi^2$  distribution. Use the fact that the cumulative  $\chi^2$  distribution can be related to the incomplete gamma function  $P(n, x)$  by (cf. Exercise 2.5)

$$F_{\chi^2}(2x; 2n) = P(n, x) \equiv \int_0^x e^{-t} t^{n-1} dt. \quad (9.5)$$

Make a sketch of  $u_\alpha(\xi)$  and  $v_\beta(\xi)$  using  $\alpha = \beta = 0.159$  and  $n = 5$ . (Quantiles of the  $\chi^2$  distribution can be looked up in standard tables or obtained from the routine `CHISIN` from the CERN program library.)

(b) Find the confidence interval  $[a, b]$  as a function of the estimate  $\hat{\xi}$ , the sample size  $n$  and the confidence levels  $\alpha$  and  $\beta$ . Suppose the estimate is  $\hat{\xi} = 1.0$ . Sketch this on the plot of  $u_\alpha(\xi)$  and  $v_\beta(\xi)$ . Evaluate  $a$  and  $b$  for  $n = 5$ ,  $\alpha = \beta = 0.159$ . Compare the result to the interval obtained from plus or minus one standard deviation about the estimate  $\hat{\xi}$ .

**Exercise 9.3:** Show that the upper and lower limits for the parameter  $p$  of a binomial distribution are

$$\begin{aligned} p_{\text{lo}} &= \frac{nF_F^{-1}[\alpha; 2n, 2(N-n+1)]}{N-n+1 + nF_F^{-1}[\alpha; 2n, 2(N-n+1)]} \\ p_{\text{up}} &= \frac{(n+1)F_F^{-1}[1-\beta; 2(n+1), 2(N-n)]}{(N-n) + (n+1)F_F^{-1}[1-\beta; 2(n+1), 2(N-n)]}. \end{aligned} \quad (9.6)$$

Here the confidence levels for the upper and lower limits are  $1 - \alpha$  and  $1 - \beta$ , respectively,  $n$  is the number of successes observed in  $N$  trials, and  $F_F^{-1}$  is the quantile of the  $F$  distribution. This is defined by the p.d.f.

$$f(x; n_1, n_2) = \binom{n_1}{n_2}^{n_1/2} \frac{\Gamma(\frac{1}{2}(n_1 + n_2))}{\Gamma(\frac{1}{2}n_1)\Gamma(\frac{1}{2}n_2)} x^{n_1/2-1} \left(1 + \frac{n_1}{n_2}x\right)^{-(n_1+n_2)/2}, \quad (9.7)$$

where  $x > 0$  and  $n_1$  and  $n_2$  are integer parameters (degrees of freedom). Use the fact that the cumulative binomial distribution is related to the cumulative distribution  $F_F(x)$  for  $n_1 = 2(n+1)$  and  $n_2 = 2(N-n)$  degrees of freedom by <sup>1</sup>

$$\sum_{k=0}^n \frac{N!}{k!(N-k)!} p^k (1-p)^{N-k} = 1 - F_F \left[ \frac{(N-n)p}{(n+1)(1-p)}; 2(n+1), 2(N-n) \right]. \quad (9.8)$$

Quantiles of the  $F$  distribution can be obtained from standard tables or computed with the routine `ffinv`. Equations (9.6) are implemented in the routines `binomlo`, `binomup` and `binomint`.

**Exercise 9.4:** In an antineutrino-nucleon scattering experiment with the Gargamelle bubble chamber at CERN, events were selected having only hadrons (from the neutral-current process  $\bar{\nu}_\mu N \rightarrow \bar{\nu}_\mu X$ ) or with hadrons and a muon (the charged-current process  $\bar{\nu}_\mu N \rightarrow \mu^+ X$ ). Out of a sample of 212 events, 64 were classified as neutral current (NC) and 148 as charged current (CC). Estimate the probability for an event to be NC and find the 68.3% central confidence

<sup>1</sup>Use of the  $F$  distribution for evaluating binomial confidence intervals is due to A. Hald, *Statistical Theory with Engineering Applications*, John Wiley, New York, 1952.

interval. Find the corresponding estimator and interval for the ratio of probabilities for NC and CC events.<sup>2</sup>

**Exercise 9.5:** In an experiment investigating particle collisions, 10 events are selected as being of a certain type, say, having a high value of some property  $x$ . Out of the 10 high- $x$  events, 2 are found to contain muons.

(a) Find the 68.3% central confidence interval for the binomial parameter  $p$  for high- $x$  events to contain muons. Express the answer as  $p = \hat{p}_{-d}^{+c}$  where  $\hat{p}$  is the ML estimate for  $p$  and  $[\hat{p} - c, \hat{p} + d]$  is the confidence interval. (The routine `binomint` can be used.)

(b) Compare the interval from (a) to  $\hat{p} \pm \hat{\sigma}_{\hat{p}}$ , where  $\hat{\sigma}_{\hat{p}}$  is the estimate of the standard deviation of  $\hat{p}$ .

(c) A common mistake is to regard the number 10 of high- $x$  events as a random variable and to include its variance in the error for  $\hat{p}$  (e.g. using error propagation). Why is this not the correct approach?

**Exercise 9.6:** Suppose that to produce the events in Exercise 9.5, the total amount of data collected corresponded to an integrated luminosity of  $L = 1 \text{ pb}^{-1}$  (known with a negligible error). The total number of events produced of a given type can be regarded as a Poisson variable  $n$  with mean value  $\nu = \sigma L$ , where  $\sigma$  is the production cross section. (Why is the Poisson distribution appropriate?)

(a) Find the 68.3% central confidence intervals for the expected numbers of events  $\nu_x$  and  $\nu_{x\mu}$  for events with high  $x$  and high  $x$  with muons, given  $n_x = 10$  and  $n_{x\mu} = 2$  events observed. What are the corresponding confidence intervals for the production cross sections,  $\sigma_x$  and  $\sigma_{x\mu}$ ?

(b) Compare the confidence intervals from (a) to intervals constructed as plus or minus one standard deviation about the estimate of the corresponding parameter.

(c) Suppose that in a separate experiment with an integrated luminosity of  $L' = 100 \text{ pb}^{-1}$ ,  $n'_x = 1173$  high- $x$  events are observed. This experiment, however, is unable to identify muons. Construct the log-likelihood function for the parameters  $\sigma_x$  and  $p = \sigma_{x\mu}/\sigma_x$  using the data  $n_x$ ,  $n_{x\mu}$  and  $n'_x$ . Show that the ML estimator for  $p$  does not depend on  $n'_x$ . Does it make sense that the result of the second experiment has no impact on the estimate of  $p$ ?

(d) Suppose the original experiment had not measured the number of high- $x$  events but had only reported the number of high- $x$  events with muons. Using only the two results  $n_{x\mu} = 2$  and  $n'_x = 1173$ , construct the log-likelihood function for  $\sigma_x$  and  $p$ . From this find the ML estimators. Use error propagation to estimate the standard deviation of  $\hat{p}$ , and compare the interval  $\hat{p} \pm \hat{\sigma}_{\hat{p}}$  to the intervals from Exercises 9.5 (a) and (b). Optional: How would you go about constructing a confidence interval for  $p$  in this case?

**Exercise 9.7:** A Particle created in an interaction is emitted at a certain angle with respect to the  $z$  axis, as shown in Fig. 9.1. A detector located a distance  $d$  from the interaction point measures the particle's position  $x$  perpendicular to the  $z$  direction. The angle  $\theta$  is defined as the angle between the  $z$  axis and the projection of the particle's trajectory into the  $(x, z)$  plane. Suppose the measured value  $x$  can be regarded as a Gaussian variable centered about the true value and having a standard deviation  $\sigma_x$ .

<sup>2</sup>In the actual experiment, small background corrections were included; see F.J. Hasert et al., Observation of neutrino-like interactions without muon or electron in the Gargamelle neutrino experiment, *Phys. Lett.* **46B** (1973) 138.

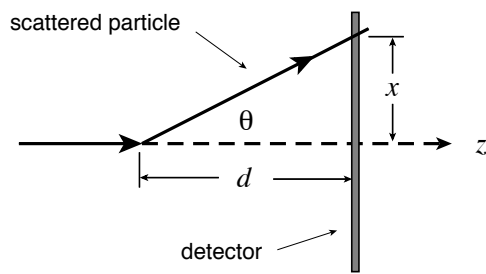


Figure 9.1: The definition of the scattering angle  $\theta$  from the trajectory projected into the  $(x, z)$  plane.

- (a) Find the central confidence interval at a confidence level  $1 - \gamma$  for the cosine of the angle  $\theta$ . Assume that the distance  $d$  is known without error.
- (b) Take  $d = 1$  m,  $\sigma_x = 1$  mm and suppose the measured value is  $x = 2.0$  mm. Find the 68.3% and 95% central confidence intervals for  $\cos \theta$ .