Statistical Data Analysis Discussion notes – week 3

- Lack of memory of the exponential distribution
- The log-normal distribution
- Generating random values from the Cauchy pdf
- Importance sampling

Example 1: "memory less ress" of exponential Exponential pdf $f(x;\xi) = \frac{-x/\xi}{\xi}$, $x \ge 0$ First, find cumulative distribution $F(x) = \int \frac{x}{5} e^{-x/5} dx' = -e^{-x/5} = 1 - e^{-x/5}$ Next, find P(x < x + x' | x > x .) C will show this is P(x x x')

Recall
$$P(A1B) = P(A \cap B)$$

 $P(B)$
For $P(x \leftarrow x_o + x' \mid x > x_o)$
 $A \leftarrow x_o + x' \mid x > x_o$
 $X_o \qquad \chi_o + x' \qquad x$
 $A \cap B = x_o \leftarrow x \leftarrow x_o + x'$
 $P(x \leftarrow x_o + x' \mid x > x_o) = \frac{P(x_o \leftarrow x \leftarrow x_o + x')}{P(x > x_o)}$

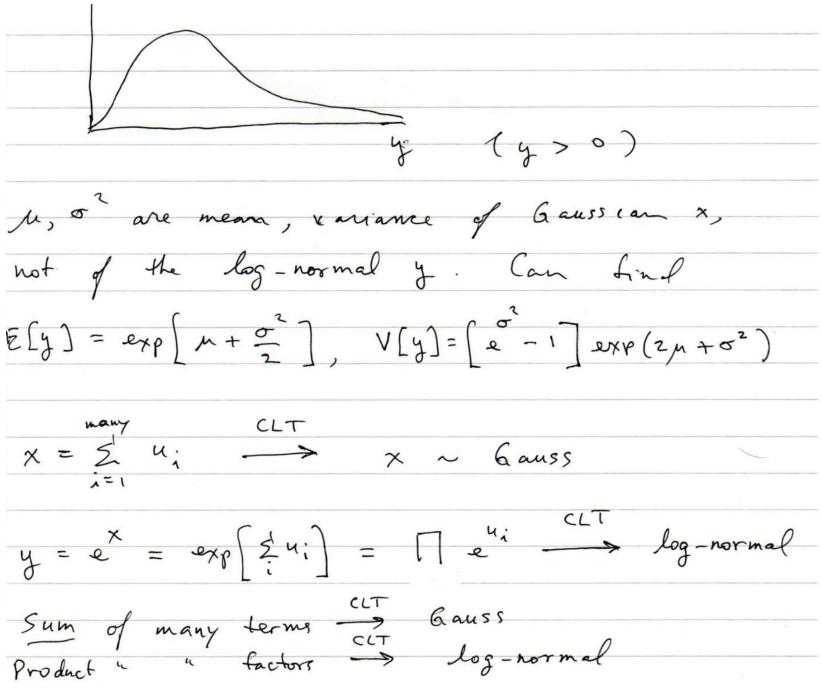
,x,+x', -x/5 5 e $F(x + x') - F(X_0)$ F(x. 100 X - ×./5 F(X) = $-(x_{o}+x')/\xi - x_{o}/\xi$ -x0/8 0 $-x'/s = F(x') = f(x \le x')$ from this using $f(x) = \frac{\partial F}{\partial x}$ $f(x-x, |x>x_{o}) = f(x)$

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memory less exponential Example cosmic ray of atmosphere top photons scintillator decays PMT created stops mu t +t mum stops, then decays decay

- (+-+.)/2 = 2.2 MS lifetime - t. Time that muon lived before to is irrelevant. Muon is just as "young" to as when it was first born: at $f(t-t_o|t>t_o) = f(t)$

Example 3 Log-normal dist. & variable trans. Gaussian $f(x; \mu, \sigma) = \frac{1}{\sqrt{2\pi}\sigma} \frac{-(x-\mu)^2/2\sigma^2}{\sqrt{2\pi}\sigma}$ y = e on final pdf of y $hy, \frac{dy}{dy} = \frac{1}{y}$ × = $f(y) = f(x(y)) \left| \frac{dx}{dy} \right| = \frac{1}{\sqrt{2\pi} \sigma y} \exp\left[-\frac{(lny)}{2}\right]$



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Random numbers from the Cauchy pdf

Cauchy pdf $f(x) = \frac{1}{\pi} \frac{1}{1+x^2}$ Cumulative dist. $F(x) = \int_{-\infty}^{\infty} \frac{dx'}{\pi(1+x'^2)}$ \rightarrow) $F(x) = \frac{1}{\pi} \tan \frac{1}{x}$ $= \frac{1}{T} \left(\tan^{-1} \times + \frac{\pi}{2} \right)$ and solve for ~ 0[0,1]

$$= \sum_{i=1}^{\infty} \chi(r) = \tan\left[\pi \left(r - \frac{1}{2}\right)\right]$$

i.e. if r_1, r_2, \dots indep. $\neq \sim \cup[0, 1]$

then $\chi_i = \chi(r_i)$ indep $\neq \sim \frac{1}{\pi(1 + \chi^2)}$

Code: cauchy MC. py

cauchy MC. ipynb

cauchyMC.py, .ipynb

```
# cauchMC.py
# simple Monte Carlo program to make histogram of uniform and Cauchy
# distributed random values and plot
# G. Cowan, RHUL Physics, updated October 2024
import matplotlib.pyplot as plt
import numpy as np
# Set random seed and other parameters
np.random.seed(12345)
numVal = 10000
nBins = 100
# Generate uniformly distributed numbers
rMin, rMax = 0., 1.
rData = np.random.uniform(rMin, rMax, numVal)
# Using transformation method, generate Cauchy distributed numbers
xMin, xMax = -10., 10.
xData = np.tan(np.pi * (rData - 0.5))
```

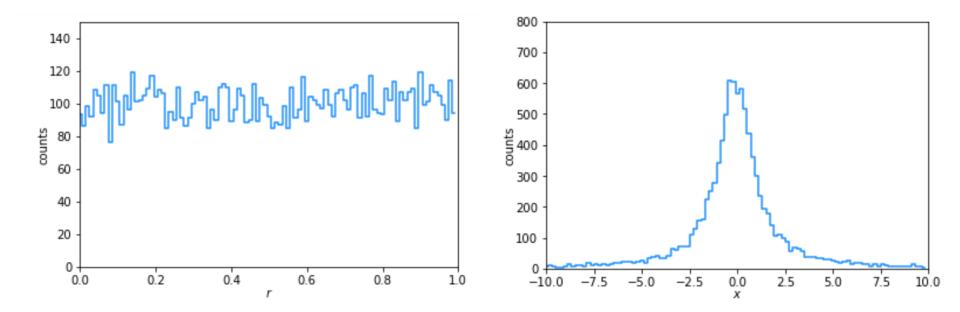
cauchyMC.py, .ipynb (continued)

```
# Define a function for plotting histograms
def plot_histogram(data, nBins, data_range, xlabel, ylabel, ylim, filename):
    fig, ax = plt.subplots()
    hist, bin_edges = np.histogram(data, bins=nBins, range=data_range)
    ax.step(bin_edges[:-1], hist, where='mid', linewidth=1.5,
            color='dodgerblue')
    ax.set xlim(data range)
    ax.set ylim(ylim)
    ax.set xlabel(xlabel, labelpad=0)
    ax.set ylabel(ylabel, labelpad=0)
    fig.subplots_adjust(bottom=0.15, left=0.15)  # Adjust layout
    plt.savefig(filename, format='pdf')
                                          # Save and show the plot
    plt.show()
# Plot and save histograms
plot histogram(rData, nBins, (rMin, rMax), r'$r$', 'counts',
               (0, 150), "uniform_histogram.pdf")
plot_histogram(xData, nBins, (xMin, xMax), r'$x$', 'counts',
               (0, 800), "cauchy histogram.pdf")
```

cauchyMC.py, .ipynb (output)

Uniform

Cauchy



Importance Sampling

Often the goal of an MC calculation is to estimate the mean value of a function h(x), where $x \sim f(x)$.

$$E_f[h(x)] = \int h(x)f(x) \, dx \equiv \mu$$

An estimator $\hat{\mu}_{MC}$ for μ is the average of N values of h(x) with x generated from f(x):

$$\hat{\mu}_{ ext{MC}} = rac{1}{N}\sum_{i=1}^{N}h(x_i)$$

This has a variance of

$$V[\hat{\mu}_{\mathrm{MC}}] = rac{1}{N} V_f[h(x)] = rac{1}{N} \left(E_f[h^2(x)] - \mu^2 \right)$$

Importance Sampling (2)

With importance sampling, one can reduce the variance for a given number of generated random values *N*.

Rewrite the desired mean value as

$$\mu = \int h(x)f(x) \, dx = \int \frac{h(x)f(x)}{g(x)} g(x) \, dx = E_g \left[\frac{h(x)f(x)}{g(x)} \right]$$

where g(x) is any other pdf nonzero on the same interval as f(x) from which we can generate random values.

Thus μ is the expectation with respect to g of h(x)f(x)/g(x), and can be estimated by generating N values of $x \sim g(x)$ and using

$$\hat{\mu}_{\mathrm{IS}} = rac{1}{N}\sum_{i=1}^{N}rac{h(x_i)f(x_i)}{g(x_i)}$$

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Importance Sampling (3)

The variance of $\hat{\mu}_{\mathrm{IS}}$ is

$$V[\hat{\mu}_{\rm IS}] = \frac{1}{N} V_g \left[\frac{h(x)f(x)}{g(x)} \right] = \frac{1}{N} \left(E_g \left[\frac{h^2(x)f^2(x)}{g^2(x)} \right] - \mu^2 \right)$$

By choosing g(x) such that h(x)f(x)/g(x) is as constant as possible, one can minimize the variance of $\hat{\mu}_{IS}$.

Minimum achieved when: $g(x) \propto |h(x)|f(x)$

Alternative estimator (smaller variance at cost of small bias):

$$\hat{\mu}_{\text{IS}} = \frac{\sum_{i=1}^{N} \frac{h(x_i)f(x_i)}{g(x_i)}}{\sum_{i=1}^{N} \frac{f(x_i)}{g(x_i)}}$$

References:

C.P. Robert and G. Casella, *Monte Carlo Statistical Methods*, 2nd ed., (Springer, New York, 2004).

J.S. Liu, Monte Carlo Strategies in Scientific Computing, (Springer, New York, 2001).

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Importance Sampling Example

We can generate uniform random numbers with the pdf:

 $f(x) = 1 , \qquad 0 \le x \le 1$

Suppose we want the mean value on [0,1] of: $h(x) = x^3 e^x$

$$\begin{split} E_f[h(x)] &= \int h(x) f(x) \, dx = \int_0^1 x^3 e^x \cdot 1 \, dx & \text{In practice we} \\ &= e^x (x^3 - 3x^2 + 6x - 6) \Big|_0^1 = 6 - 2e = 0.563436 \end{split}$$

We need a pdf g(x) from which we can sample such that

$$\frac{h(x)f(x)}{g(x)}$$

is approximately constant on the relevant interval 0 < x < 1.

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Importance Sampling Example (2)

Guess: $g(x) = (\theta + 1)x^{\theta}$ for some θ that we can adjust.

The cumulative distribution is G

$$G(x)=\int_0^x g(x')\,dx'=x^{ heta+1}$$

To sample from g(x), set g(x) = rwhere $r \sim U[0,1]$ and solve for x:

For e.g. $\theta = 3.4$ and N = 1000,

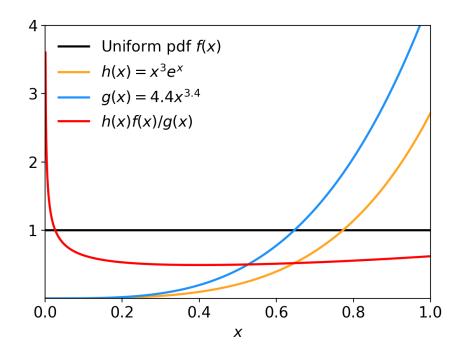
Exact value = 0.563436

Monte Carlo estimate: 0.525473 +/- 0.022596

Importance sampling estimate: 0.564376 +/- 0.001159

20 times smaller err. than MC

$$x(r) = r^{1/(\theta+1)}$$



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