Statistical Data Analysis Problem Sheet 9 G. Cowan / December 2020

1(a) [3 marks] As the x_i are all independent, the likelihood is found from the product of pdfs,

$$L(v) = \prod_{i=1}^{N} \frac{1}{\sqrt{2\pi v}} e^{-x_i^2/2v} ,$$

and the log-likelihood is

$$\ln L(v) = -\frac{N}{2} \ln v - \frac{1}{2v} \sum_{i=1}^{N} x_i^2 + C ,$$

where C represents terms that do not depend on v. Setting the derivative of $\ln L$ to zero,

$$\frac{\partial \ln L}{\partial v} = -\frac{N}{2v} + \frac{1}{2v^2} \sum_{i=1}^N x_i^2 = 0 ,$$

and solving for v gives the ML estimators,

$$\hat{v} = \frac{1}{N} \sum_{i=1}^{N} x_i^2 \; .$$

1(b) [2 marks] Using $E[x_i] = 0$ and therefore $E[x_i^2] = v$, the expectation value of \hat{v} is

$$E[\hat{v}] = \frac{1}{N} \sum_{i=1}^{N} E[x_i^2] = v ,$$

and therefore the bias is zero.

1(c) [3 marks] We are told that $t = N\hat{v}/v$ follows a chi-squared distribution for N degrees of freedom, and therefore has variance 2N. From the variance of t we have

$$V[t] = V\left[\frac{N\hat{v}}{v}\right] = \frac{N^2}{v^2}V[\hat{v}] ,$$

and since V[t] = 2N we find

$$V[\hat{v}] = \frac{2v^2}{N} \; .$$

1(d) [3] marks] Again using $t = N\hat{v}/v$, we are told to take small values of t as representing lower compatibility. The p-value of v is therefore

$$p_v = P(t \le t_{\text{obs}} | v) = F_{\chi^2_N}(t_{\text{obs}})$$

where $F_{\chi_N^2}$ is the chi-squared cumulative distribution for N degrees of freedom, and where $t_{\rm obs}$ refers to the observed value of the statistic. Taking $\hat{v} = v t_{\rm obs}/N$ to mean the observed value of the estimator, we obtain the desired result,

$$p_v = F_{\chi_N^2}(N\hat{v}/v) \; .$$

1(e) [3 marks] To find the upper limit on v, we set the *p*-value equal to $\alpha = 1 - CL$ and solve for v, i.e.,

$$p_v = F_{\chi^2_N}(N\hat{v}/v) = \alpha \; .$$

Applying the inverse of the cumulative distribution (the chi-squared quantile) to each side gives

$$\frac{N\hat{v}}{v} = F_{\chi_N^2}^{-1}(\alpha) \; .$$

Solving for v gives the upper limit,

$$v_{\rm up} = \frac{N\hat{v}}{F_{\chi_N^2}^{-1}(\alpha)} \ .$$

1(f) [3 marks] The Jeffreys prior is $\pi_J(v) \propto \sqrt{I}$ where I is the Fisher information. For this we need

$$\frac{\partial^2 \ln L}{\partial v^2} = \frac{N}{2v^2} - \frac{1}{v^3} \sum_{i=1}^N x_i^2 \ . \label{eq:eq:alpha_eq}$$

The Fisher information is therefore

$$I = -E\left[\frac{\partial^2 \ln L}{\partial v^2}\right] = -\frac{N}{2v^2} + \frac{1}{v^3} \sum_{i=1}^{N} E[x_i^2] = \frac{N}{2v^2} ,$$

where to obtain the final equality we used $E[x_i^2] = v$. Therefore the Jeffreys prior is

$$\pi_{\rm J}(v) \propto \sqrt{I} \propto rac{1}{v} \; .$$

1(g) [3 marks] Using Bayes' theorem as a proportionality, the posterior probability for v given $\vec{x} = (x_1, \ldots, x_N)$ is

$$p(v|\vec{x}) \propto p(\vec{x}|v) \pi_{\rm J}(v) \propto \frac{1}{v} \prod_{i=1}^{N} \frac{1}{\sqrt{2\pi v}} e^{-x_i^2/2v} \propto v^{-N/2-1} \exp\left[-\frac{1}{2v} \sum_{i=1}^{N} x_i^2\right]$$

To find the posterior mode we set its derivative to zero. Letting $Q = \sum_{i=1}^{N} x_i^2$ we obtain

$$\frac{\partial p(v|\vec{x})}{\partial v} \propto v^{-N/2-1} e^Q \frac{Q}{2v^2} + e^Q \left(-\frac{N}{2} - 1\right) v^{-N/2-2} = 0 \; .$$

Solving for v gives the mode,

$$mode[v] = \frac{1}{N+2} \sum_{i=1}^{N} x_i^2 .$$